# The Ordered and Multinomial Models 

Quantitative Microeconomics

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## Outline

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## Motivation

We consider the following two extensions from binary dependent models:

- Ordered response models: The dependent variable takes a number of finite and discrete values that contain ordinal information.
- Multinomial response models: The dependent variable takes a number of finite and discrete values that DO NOT contain ordinal information.

As in the probit and logit cases, the dependent variable is not strictly continuous. Estimation will be carried out using the ML estimator.

## Examples of ordered models

- Credit rating, using seven categories, from "absolutely not credit worthy" to "credit worthy".
- Decision to remain inactive, to work part-time, or to work full-time.
- In an income regression, income levels are coded in intervals: $[0,1000),[1000,1500)[1500,2000),[2000, \infty)$
- On value statements, several answers with ordinal content: "completely disagree", "disagree", "somewhat agree", "completely agree"


## Examples of multinomial models

- Choice of transport mode: train, bus, car
- Economic status: inactive, unemployed, self-employed, employee
- Education field choice: hard science, health sciences, social sciences, humanities


## Ordered Response Models

- The two standard models are the ordered probit and the ordered logit.
- The approach is equivalent: we simply use for the ordered probit the normal CDF $\Phi()$ and for the ordered logit the logistic CDf $\Lambda()$.
- OLS does not work because the dependent variable does not have cardinal meaning:
- credit worthiness: $0,1,2,3,4,5$ : the change from 0 to 1 does not have to be "equivalent" to the change from 4 to 5 .
- activity: inactive $=0$, part-time $=1$, full-time $=2$ : While inactive is zero hours of work, in practice code 1 reflects any hours of work between 1 and (usually) 30 hours of work, and code 2 reflects more 30 hours of work. This implies that there is no proportionality in going from 0 to 1 and going from 1 to 2 .


## Simplification

- Binary choice models (LPM, probit, logit) could potentially be used by grouping all categories into two major ones,
- This is the case when the sample is small and the ordinal categories can be logically be grouped in two major categories.
- In some cases, this is probably a very bad idea (income intervals).
- Consider three observed outcomes: $y=0,1,2$.
- Consider the latent variable model without a constant:

$$
y^{*}=\beta_{1} x_{1}+\ldots+\beta_{k} x_{k}+\varepsilon
$$

where $\varepsilon \sim \mathscr{N}(0,1)$.

- Define two cut-off points: $\alpha_{1}<\alpha_{2}$
- We do not observe $y^{*}$, but we observe choices according to the following rule

$$
\begin{aligned}
& y=0 \text { if } y^{*} \leq \alpha_{1} \\
& y=1 \text { if } \alpha_{1}<y^{*} \leq \alpha_{2} \\
& y=2 \text { if } \alpha_{2}<y^{*}
\end{aligned}
$$

## Example: activity

- $y=0$ if inactive, $y=1$ if part-time, $y=2$ if full-time
- $y^{*}=\beta_{e} \times$ educ $+\beta_{k} \times k i d s+\varepsilon$, where $\varepsilon \sim \mathscr{N}(0,1)$
- Then

$$
\begin{aligned}
& y=0 \text { if } \beta_{e} \times e d u c+\beta_{k} \times k i d s+\varepsilon \leq \alpha_{1} \\
& y=1 \text { if } \alpha_{1}<\beta_{e} \times e d u c+\beta_{k} \times k i d s+\varepsilon \leq \alpha_{2} \\
& y=2 \text { if } \alpha_{2}<\beta_{e} \times e d u c+\beta_{k} \times k i d s+\varepsilon
\end{aligned}
$$

- Note that we could alternatively introduce a constant $\beta_{0}$ and assume that $\alpha_{1}=0$.


## Interpretation

- As in other nonlinear models, marginal effects can be computed to learn about the partial effects of a small change in explanatory variable $x_{j}$.
- For ordered models we can compute marginal effects on the predicted probabilities along the same principles.


## Partial effects on predicted probabilities

- For binary choice models, we focused on the effects on the probability that $y$ is equal to one.
- In the ordered models, things are not so simple: we now have more than two outcomes:

$$
\begin{aligned}
& \frac{\partial \operatorname{Pr}(y=0 \mid x)}{\partial x_{j}}=-\phi\left(x^{\prime} \beta-\alpha_{1}\right) \beta_{j} \\
& \frac{\partial \operatorname{Pr}(y=1 \mid x)}{\partial x_{j}}=\left(\phi\left(x^{\prime} \beta-\alpha_{1}\right)-\phi\left(x^{\prime} \beta-\alpha_{2}\right)\right) \beta_{j} \\
& \frac{\partial \operatorname{Pr}(y=2 \mid x)}{\partial x_{j}}=\phi\left(x^{\prime} \beta-\alpha_{2}\right) \beta_{j}
\end{aligned}
$$

- if $x_{j}$ is discrete we compute as in the binary case the discrete change in the predicted probabilities associated with changing $x_{j}$.


## Partial Effects

- The partial effect of $x_{j}$ on the predicted probability of
- the highest outcome has the same sign as $\beta_{j}$.
- the lowest outcome has the opposite sign to $\beta_{j}$
- intermediate outcomes cannot, in general, be inferred from the sign of $\beta_{j}$.
- The last results is due to two offsetting effects. Suppose $\beta_{j}>0$ and you increase $x_{j}$. The intermediate category
- may become more likely since the probability of the lowest category falls.
- may also become less likely because the probability of the highest category increases.
- Typically, partial effects for intermediate probabilities are quantitatively small and often statistically insignificant.


## Discussion

How best to interpret results from ordered models?

- One option is to look at the estimated $\beta$-parameters, emphasizing the underlying latent variable equation with which we started.
- Another option might be to look at the effect on the expected value of the ordered response variable, e.g.
$\frac{\partial \mathrm{E}(y \mid x)}{\partial x_{j}}=\frac{\partial \operatorname{Pr}(y=0 \mid x)}{\partial x_{j}} \times 0+\frac{\partial \operatorname{Pr}(y=1 \mid x)}{\partial x_{j}} \times 1+\frac{\partial \operatorname{Pr}(y=2 \mid x)}{\partial x_{j}} \times 2$
This may make a lot of sense if $y$ is a numerical variable, as in the income variable.
- Alternatively, you might just want to report the effect on the probability of observing the ordered categories.


## Multinomial Response

- The dependent variable is such that
- more than two outcomes are possible
- the outcomes cannot be ordered in any natural way.
- Again, we could bunch two or more categories and so construct a binary outcome variable from the raw data, but in doing so, we throw away potentially interesting information.
- OLS is also not a good model in this context.
- However, the logit model for binary choice can be extended to model more than two outcomes.


## Random Utility Model

- Assume that there are three transport alternatives: bus, car, train:

$$
\begin{aligned}
& U_{b}=x_{b}^{\prime} \beta_{b}+\varepsilon_{b} \\
& U_{c}=x_{c}^{\prime} \beta_{c}+\varepsilon_{c} \\
& U_{t}=x_{t}^{\prime} \beta_{t}+\varepsilon_{t}
\end{aligned}
$$

where $\left\{\varepsilon_{b}, \varepsilon_{c}, \varepsilon_{t}\right\}$ are the effects on utility unobserved by the econometrician

$$
\begin{aligned}
& \text { If } x_{b}^{\prime} \beta_{b}+\varepsilon_{b} \geq \max \left\{x_{c}^{\prime} \beta_{c}+\varepsilon_{c}, x_{t}^{\prime} \beta_{t}+\varepsilon_{t}\right\} \text { then } y=0 \\
& \text { If } x_{c}^{\prime} \beta_{c}+\varepsilon_{c}>\max \left\{x_{b}^{\prime} \beta_{b}+\varepsilon_{b}, x_{t}^{\prime} \beta_{t}+\varepsilon_{t}\right\} \text { then } y=1 \\
& \text { If } x_{t}^{\prime} \beta_{t}+\varepsilon_{t}>\max \left\{x_{c}^{\prime} \beta_{c}+\varepsilon_{c}, x_{t}^{\prime} \beta_{t}+\varepsilon_{t}\right\} \text { then } y=2
\end{aligned}
$$

## Notation

- We have two unobserved independent effects

$$
\begin{aligned}
& \varepsilon_{01}=\varepsilon_{b}-\varepsilon_{c} \\
& \varepsilon_{02}=\varepsilon_{b}-\varepsilon_{t}
\end{aligned}
$$

- note that $\varepsilon_{12}=\varepsilon_{c}-\varepsilon_{t}=\varepsilon_{02}-\varepsilon_{01}$
- Define

$$
\begin{aligned}
x_{b}^{\prime} \beta_{b}-x_{c}^{\prime} \beta_{c} & =x^{\prime} \beta_{01} \\
x_{b}^{\prime} \beta_{b}-x_{t}^{\prime} \beta_{t} & =x^{\prime} \beta_{02}
\end{aligned}
$$

## Assumption

$$
\left\{\varepsilon_{01}, \varepsilon_{02}\right\} \sim F
$$

where $F$ is symmetric.
Then

$$
\begin{aligned}
\operatorname{Pr}(y=0 \mid x) & =\operatorname{Pr}\left(x^{\prime} \beta_{01}+\varepsilon_{01} \geq 0, x^{\prime} \beta_{02}+\varepsilon_{02} \geq 0 \mid x\right) \\
& =\operatorname{Pr}\left(\varepsilon_{01} \geq-\left(x^{\prime} \beta_{01}\right), \varepsilon_{02} \geq-\left(x^{\prime} \beta_{02}\right) \mid x\right)
\end{aligned}
$$

Given symmetry,

$$
\operatorname{Pr}(y=0 \mid x)=F\left(x^{\prime} \beta_{01}, x^{\prime} \beta_{02}\right)
$$

## Multinomial Logit

- We must model the probability that an individual belongs to category $j$ conditional to having characteristics $x$ :

$$
\operatorname{Pr}(y=j \mid x)
$$

- When vector $\left\{\varepsilon_{b}, \varepsilon_{c}, \varepsilon_{t}\right\}$ has a extreme value distribution, then we have the Multinomial Logit:

$$
\begin{aligned}
& \operatorname{Pr}(y=0 \mid x)=1-\operatorname{Pr}(y=1 \mid x)-\operatorname{Pr}(y=2 \mid x) \\
& \operatorname{Pr}(y=1 \mid x)=\frac{\exp \left(x^{\prime} \beta_{1}\right)}{1+\exp \left(x^{\prime} \beta_{1}\right)+\exp \left(x^{\prime} \beta_{2}\right)} \\
& \operatorname{Pr}(y=2 \mid x)=\frac{\exp \left(x^{\prime} \beta_{2}\right)}{1+\exp \left(x^{\prime} \beta_{1}\right)+\exp \left(x^{\prime} \beta_{2}\right)}
\end{aligned}
$$

- The main difference compared to the binary logit is that there are now two parameter vectors, $\beta_{1}$ and $\beta_{2}$
- in the general case with $J$ possible responses, there are $J-1$ parameter vectors.
- This makes interpretation of the coefficients more difficult than for binary choice models.


## Interpretation with three alternatives

The easiest case to think about is where $\beta_{1 j}$ and $\beta_{2 j}$ have the same sign.

- If $\beta_{1 j}$ and $\beta_{2 j}$ are positive then an increase in the variable $x_{j}$ it less likely that the individual belongs to category $0 .$. .
- and $\operatorname{Pr}\left(y_{i}=1 \mid x_{i}\right)+\operatorname{Pr}\left(y_{i}=2 \mid x_{i}\right)$ increases
- to know how this total increase is allocated between these two probabilities, we need to look at the marginal effects: the partial derivative is very complex and the marginal effect $\frac{\partial \operatorname{Pr}(y=1 \mid x)}{\partial x_{j}}$ may in fact be negative even if $\beta_{1 j}$ !


## Independence of irrelevant alternatives (IIA)

- One important limitation of the multinomial logit is that the ratio of any two probabilities $/$ and $m$ depends only on the parameter vectors $\beta_{l}$ and $\beta_{m}$ and the explanatory variables $x$

$$
\begin{aligned}
\frac{\operatorname{Pr}(y=1 \mid x)}{\operatorname{Pr}(y=2 \mid x)} & =\frac{\exp \left(x^{\prime} \beta_{1}\right)}{\exp \left(x^{\prime} \beta_{2}\right)} \\
& =\exp \left(x^{\prime}\left(\beta_{1}-\beta_{2}\right)\right)
\end{aligned}
$$

- The inclusion or exclusion of other categories is irrelevant to the ratio of the two probabilities.
- This behavior is referred to as the "independence of irrelevant alternatives", and it can lead to counter-intuitive behavior


## Example: IIA can be counter-intuitive

- Individuals can commute to work by three transportation means: blue bus, red bus, or train.
- Individuals choose one of these alternatives, and the econometrician estimates a multinomial logit modeling this decision, and obtains an estimate of

$$
\frac{\operatorname{Pr}(y=\text { red } \mid x)}{\operatorname{Pr}(y=\operatorname{train} \mid x)}=\exp \left(x^{\prime}\left(\beta_{r e d}-\beta_{\text {train }}\right)\right)
$$

- Suppose that the bus company now removes the blue bus from the set of options, do you think that $\frac{\operatorname{Pr}(y=r e d \mid x)}{\operatorname{Pr}(y=\operatorname{train} \mid x)}$ would be the same as before?


## Other multinomial models

- There are lots of other econometric models that can be used to model multinomial response models:
- multinomial probit,
- conditional logit,
- nested logit
- They are beyond the scope of the course.


## Summary

- When the dependent variable has a finite number of discrete values, we can extend the probit and logit models
- When the dependent variable entails some ordinal information, then we can use ordered probit and logit models
- When the dependent variable does not contain any ordinal information, we can use multinomial models. One such example is the multinomial logit.
- These are all nonlinear models, and they can all be estimated by MLE.

